Exam 2 for EECS3030(02), Spring 2020

10:10-11:50 Weddnesday, June 17, 2020

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(30pts) 1. Fill the following blanks.

(a) Let $\{X_i \sim b(16, 0.75), 1 \le i \le 10\}$ be a random sample from a binomial distribution. Define $X = \sum_{i=1}^{10} X_i$, then

the moment-generating function $M_X(t) = \underline{(0.25 + 0.75e^t)^{160}}, \quad Var(X) = \underline{30}$ the probability mass function of X, $f_X(x) = \underline{C(160, x)(0.75)^x(0.25)^{160-x}}, \quad 0 \le x \le 160$

(b) Let $\{Y_i, 1 \leq i \leq 8\}$ be a random sample of Poisson distribution with variance $Var(Y_8) = 2$ and define $Y = \sum_{i=1}^{8} Y_i$. Then

the moment-generating function $M_Y(t) = e^{16(e^t-1)}$

the p.m.f. of
$$Y$$
, $f_Y(y) = \underbrace{\frac{e^{-16}(16)^y}{y!}}, y = 0, 1, \dots, Var(Y) = \underline{16}$

- (c) Let the joint probability distribution function of the lifetimes of two brands of mobile phones be given by $F(x,y)=(1-e^{-x})(1-e^{-y})$ if $x>0,\ y>0$. Then the probability that one mobile phone lasts more than three times as long as the other is $\frac{1}{2}$
- (d) Let the joint probability density function of random variables X and Y be given by $f(x,y) = 8xy, 0 \le x \le y \le 1$. Then

$$f_X(x) = 4x - 4x^3, \ 0 \le x \le 1, \ E(Y) = \frac{4}{5}, \ \rho(X, Y) = \frac{4}{\sqrt{66}}$$

(e) Let the conditional probability density function of X, given that Y = y, be

$$f_{X|Y}(x|y) = \frac{x+y}{1+y}e^{-x}, \quad 0 < x < \infty, \ 0 < y < \infty.$$

$$P(X < 4|Y = 1) = 1 - 3e^{-4}$$

(30pts) 2. Fill the following blanks.

(a) Suppose that X and Y are independent and identically distributed exponential random variables with variance 4. Define $W = \frac{Y}{X+Y}$. Then

the probability density function $f_Y(y) = \frac{1}{2}e^{-y/2}, y \ge 0$

$$E(W) = \frac{1}{2}, \quad Var(W) = \frac{1}{12}$$

(b) Let $\{X_i, 1 \le i \le 12\}$ be a random sample with the probability density function $f_{X_1}(x) = \frac{1}{16}x^2e^{-x/2}, x > 0$. Define $W = \sum_{i=1}^{12} X_i$. Then

the moment-generating function $M_{X_3}(t) = \frac{1}{(1-2t)^3}, \ Var(X_3) = \underline{12}$

the moment-generating function $M_W(t) = \frac{1}{(1-2t)^{36}}$

(c) Let $\{X_1, X_2, \dots, X_9\}$ be a random sample of size n=9 from the exponential distribution with $Var(X_4) = 4$. Define $Y = \sum_{i=1}^{9} X_i$. Then

the moment-generating function of Y, $\phi_Y(t) = \frac{1}{(1-2t)^9}$,

the p.d.f. of
$$Y$$
, $f_Y(y) = \frac{1}{\Gamma(9)2^9} y^8 e^{-y/2}$, $y > 0$, $Var(Y) = \underline{36}$

(d) Let $Z \sim N(0,1)$. Define Y = 3Z + 2, then

the p.d.f. of
$$Y$$
, $f_Y(y) = \frac{1}{\sqrt{18\pi}} e^{-(y-2)^2/18}, -\infty < y < \infty$

the moment-generating function of Y, $\phi_Y(t) = \underline{e^{2t + \frac{9t^2}{2}}}$, $Var(Y) = \underline{9}$

(e) Let $\{X_i \sim b(12, 0.5), 1 \leq i \leq n\}$ be a random sample from a binomial distribution of size n. Define $W_n = (\sum_{i=1}^n X_i - 6n)/\sqrt{3n}$.

 $E(X_n) = \underline{}_{0}, \quad Var(X_n) = \underline{}_{0}$ According to the Central Limit Theorem,

the limiting moment-generating function, $limit_{n\to\infty}M_{W_n}(t)=\underline{e^{t^2/2}}$

- (10pts) 3. Let $\{X_i \sim N(3,4), 1 \leq i \leq n\}$ be a random sample of a normal distribution. Define $W = \sum_{i=1}^n \frac{X_i 3}{2}, Y = \sum_{i=1}^n \left(\frac{X_i 3}{2}\right)^2$.
 - (a) Write down the probability density function of X_n .
 - (b) Calculate the moment-generating function of W and its probability density function, $f_W(w)$, respectively.
 - (c) Calculate the moment-generating function of Y and its probability density function, $f_Y(y)$, respectively.
 - (d) Calculate Var(W), Var(Y), respectively.

(Sa)
$$f_{X_n}(x) = \frac{1}{\sqrt{8\pi}} e^{-(x-3)^2/8}, -\infty < x < \infty$$

- (Sb) Let $Z_i = \frac{X_i 3}{2} \sim N(0, 1)$, then $W = \sum_{i=1}^n Z_i$. Hence, $M_W(t) = e^{nt^2/2}$, and $f_W(w) = \frac{1}{\sqrt{2n\pi}}e^{-w^2/2n}$, $-\infty < w < \infty$
- (Sc) $Y \sim \chi^2(n)$, then $M_Y(t) = \frac{1}{(1-2t)^{n/2}}$, $f_Y(y) = \frac{1}{\Gamma(\frac{n}{2})2^{n/2}}y^{n/2-1}e^{-y/2}$, y > 0.
- (Sd) Var(W) = n, Var(Y) = 2n.

(10 pts) 4. The time it takes for a student to finish an aptitude test (in hours) has the probability density function

$$f(x) = 6(x-1)(2-x)$$
 if $1 < x < 2$, and 0 elsewhere.

Approximate the probability that the average length of time it takes for a random sample of 16 students to complete the test in less than 1 hour and 45 minutes (that is, 1.75 hours).

(**Hint**) You can express your solution in terms of $\Phi(r)$, where $\Phi(x) = \int_{-\infty}^{x} \frac{1}{\sqrt{2\pi}} e^{-t^2/2} dt$, and $r \in (-4.99, 4.99)$.

(Solution)

$$E(X) = \int_{1}^{2} x \cdot 6(x - 1)2 - x) dx = \int_{1}^{2} -6x^{3} + 18x^{2} - 12x dx = \frac{3}{2},$$

$$E(X^{2}) = \int_{1}^{2} x^{2} \cdot 6(x - 1)2 - x) dx = \int_{1}^{2} -6x^{4} + 18x^{3} - 12x^{2} dx = \frac{23}{10},$$

$$Var(X) = \frac{23}{10} - (\frac{3}{2})^{2} = \frac{1}{20}.$$

Let $\bar{X} = \sum_{i=1}^{16} X_i$, by Central Limit Theorem, $P\left(\frac{\bar{X} - \frac{3}{2}}{\frac{1}{\sqrt{20}} \cdot \frac{1}{\sqrt{16}}} \le \frac{1.75 - \frac{3}{2}}{\frac{1}{8\sqrt{5}}}\right) \approx \Phi(2\sqrt{5}) = \Phi(4.472)$.

- (10pts) 5. Let $\{X_1, X_2, X_3, X_4\}$ be a random sample from the uniform distribution U(0,1) with the set of order statistics $X_{(1)} < X_{(2)} < X_{(3)} < X_{(4)}$.
 - (a) Calculate the distribution function of $X_{(3)}$, $F_{X_{(3)}}(x) = P(X_{(3)} \le x)$.
 - (b) Calculate the probability density function of $f_{X_{(3)}}(x)$.
 - (c) Calculate $E[X_{(3)}]$.

(Sa)
$$F(x) = F_{X_1}(x) = x$$
, $0 \le x \le 1$, then
$$F_{X_{(3)}}(x) = P(X_{(3)} \le x)$$

$$= \sum_{k=3}^{4} \binom{a}{b} [F(x)]^k [1 - F(x)]^{4-k}$$

$$= 4x^3 (1-x) + x^4 = 4x^3 - 3x^4, \quad 0 \le x \le 1$$

(Sb)
$$f_{X_{(3)}}(x) = 12x^2 - 12x^3, \quad 0 \le x \le 1.$$

(Sc)
$$E[X_{(3)}] = \int_0^1 x f_{X_{(3)}}(x) dx = \int_0^1 12x^3 - 12x^4 dx = \frac{3}{5}$$
.

- (10pts) 6. While rolling a balanced die of six faces successively, the first 3 occurred on the third roll. What is the expected number of rolls until the first 1 appears?
- (Solution) Let X, Y be the number of rolls until the first 1 appears and the first 3 occurs, respectively. We want to calculate

$$E(X|Y=3) = \sum_{x=1}^{\infty} x p_{X|Y}(x|y=3) = \sum_{x=1}^{\infty} x \frac{p(x,3)}{p_Y(3)},$$

where $p_Y(3) = (\frac{5}{6})^2(\frac{1}{6}) = \frac{25}{216}$, $p(1,3) = (\frac{1}{6})(\frac{5}{6})(\frac{1}{6}) = \frac{5}{216}$, $p(2,3) = (\frac{4}{6})(\frac{1}{6})(\frac{1}{6}) = \frac{4}{216}$, p(3,3) = 0, and for x > 3,

$$p(x,3) = (\frac{4}{6})^2 (\frac{1}{6})(\frac{5}{6})^{x-4} (\frac{1}{6}) = \frac{1}{81}(\frac{5}{6})^{x-4}.$$

Therefore,

$$E(X|Y=3) = \sum_{x=1}^{\infty} x p_{X|Y}(x|y=3) = \sum_{x=1}^{\infty} x \frac{p(x,3)}{p_Y(3)} = \frac{157}{25} = 6\frac{7}{25}.$$