- ► Solution of assignment 3.
- ► Hint of assignment 4.

- ► Solution of assignment 3.
- ► Hint of assignment 4.

Assignment 3 - problem 1

We prove that if Z is a Poisson random variable of mean µ, where µ ≥ 1 is an integer, then Pr(Z ≥ µ) ≥ 1/2.

1. Show that
$$Pr(Z = \mu + h) \ge Pr(Z = \mu - h - 1)$$
 for
 $0 \le h \le \mu - 1$.
Ans. For any non-negative integer k ,
 $Pr(Z = k) = e^{-\mu} \mu^k / k!$ by definition. So,

$$\frac{\Pr(Z = \mu + h)}{\Pr(Z = \mu - h - 1)} = \frac{\mu^{2h+1}}{(\mu - h)(\mu - (h - 1))\cdots(\mu + h)}$$
$$= \frac{\mu^2}{\mu^2 - h^2} \cdot \frac{\mu^2}{\mu^2 - (h - 1)^2} \cdots \frac{\mu^2}{\mu^2 - 1^2} \cdot \frac{\mu}{\mu}$$
$$\geq 1.$$

In other words, $\Pr(Z = \mu + h) \ge \Pr(Z = \mu - h - 1).$

Assignment 3 - problem 1 (con't)

▶ Using part (a), argue that $Pr(Z \ge \mu) \ge 1/2$. Ans.

$$1 = \sum_{k=1}^{\infty} \Pr(Z = k) = \sum_{k=0}^{\mu-1} \Pr(Z = k) + \Pr(Z \ge \mu).$$

Then by part (a), we know that

$$\sum_{k=0}^{\mu-1} \Pr(Z=k) \leq \sum_{k=\mu}^{2\mu-1} \Pr(Z=k) \leq \sum_{k=\mu}^{\infty} \Pr(Z=k) = \Pr(Z \geq \mu).$$

Combining, we have

 $1 = \sum_{k=0}^{\infty} \Pr(Z = k) = \sum_{k=0}^{\mu-1} \Pr(Z = k) + \Pr(Z \ge \mu) \le 2 \Pr(Z \ge \mu).$

Thus, $\Pr(Z \ge \mu) \ge 1/2$.

 In Page 15 of Lecture Notes 14, we showed that for any nonnegatvie function f,

$$E\left[f(Y_1^{(m)},\cdots,Y_n^{(m)})\right] \ge E\left[f(X_1^{(m)},\cdots,X_n^{(m)})\right] \Pr\left(\sum Y_i^{(m)}=m\right)$$

1. Now, suppose we further know that $E\left[f(X_1^{(m)}, \cdots, X_n^{(m)})\right]$ is monotonically increasing in *m*. Show that

$$E\left[f(Y_1^{(m)},\cdots,Y_n^{(m)})\right] \ge E\left[f(X_1^{(m)},\cdots,X_n^{(m)})\right] \Pr\left(\sum Y_i^{(m)} \ge m\right)$$

Assignment 3 - problem 2 (con't)

► Proof.

$$E\left[f(Y_1^{(m)}, \dots, Y_n^{(m)})\right]$$

$$= \sum_{k\geq 0} E\left[f(Y_1^{(m)}, \dots, Y_n^{(m)}) \mid \sum Y_i^{(m)} = k\right] \Pr\left(\sum Y_i^{(m)} = k\right)$$

$$\geq \sum_{k\geq m} E\left[f(Y_1^{(m)}, \dots, Y_n^{(m)}) \mid \sum Y_i^{(m)} = k\right] \Pr\left(\sum Y_i^{(m)} = k\right)$$

$$= \sum_{k\geq m} E\left[f(X_1^{(k)}, \dots, X_n^{(k)})\right] \Pr\left(\sum Y_i^{(m)} = k\right)$$

$$\geq \sum_{k\geq m} E\left[f(X_1^{(m)}, \dots, X_n^{(m)})\right] \Pr\left(\sum Y_i^{(m)} = k\right)$$

monotonically increasing.

$$= E\left[f(X_1^{(m)}, \cdots, X_n^{(m)})\right] \Pr\left(\sum Y_i^{(m)} \ge m\right),$$

Assignment 3 - problem 2 (con't)

Combining part (a) with the results in Question 1, prove the monotonically increasing case of theorem in Page 20 of Lecture Notes 14.

Proof. Let $Z = \sum Y_i^{(m)}$. Since each $Y_i^{(m)}$ is a Poisson random variable, their sum Z is also a Poisson random variable. Further, the mean value of Z is m. Thus, by results in Question 1, $\Pr(Z \ge m) \ge 1/2$. Combining this with part (a), we have

$$E\left[f(X_1^{(m)},\cdots,X_n^{(m)})\right] \leq 2E\left[f(Y_1^{(m)},\cdots,Y_n^{(m)})\right]$$

when $E\left[f(X_1^{(m)}, \dots, X_n^{(m)})\right]$ is monotonically increasing in m. This completes the proof of the monotonically increasing case of the desired theorem.

Bloom filters can be used to estimate set differences. Suppose you have a set X and I have a set Y, both with n elements. For example, the sets might represent our 100 favorite songs. We both create Bloom filters of our sets, using the same number of bits m and the same k hash functions. Determine the expected number of bits where our Bloom filters differ as a function of m, n, k, and |X ∩ Y|.

- ► Ans. Let Z be a random variable denoting the number of bits where the Bloom filters differ. Let Z_i be an indicator such that
 - $Z_i = 1$ if the *i*th bit of the Bloom filters differ $Z_i = 0$ otherwise.

Thus, $Z = Z_1 + Z_2 + \cdots + Z_m$

- When |X ∩ Y| = r, Z_i = 1 only happens when each of the r common elements are not mapped to the *i*th bit, together with exactly one of the following cases (that causes the *i*th bit different):
 - 1. Some elements of $X (X \cap Y)$ is mapped to the *i*th bit, but all elements of $Y (X \cap Y)$ are not;
 - 2. Some elements of $Y (X \cap Y)$ is mapped to the *i*th bit, but all elements of $X (X \cap Y)$ are not.

Let Q_i denote the event that the r common elements are not mapped to the *i*th bit. By assuming that the hash functions we choose will map elements independently and uniformly at random to one of the m bits, we have

$$Pr(Z_i = 1) = Pr(Q_i \cap (Case (a) \text{ or } Case (b)))$$

= $Pr(Q_i) (Pr(Case (a)) + Pr(Case (b)))$
= $\left(1 - \frac{1}{m}\right)^{rk} (Pr(Case (a)) + Pr(Case (b)))$

Assignment 3 - problem 3 (con't)

$$= \left(1 - \frac{1}{m}\right)^{rk} \times 2 \times \underbrace{\left(1 - \left(1 - \frac{1}{m}\right)^{(n-r)k}\right)}_{\text{some elements mapped to ith bit}}$$
$$\times \underbrace{\left(\left(1 - \frac{1}{m}\right)^{(n-r)k}\right)}_{\text{no elements mapped to ith bit}}$$
$$= 2\left(1 - \frac{1}{m}\right)^{nk}\left(1 - \left(1 - \frac{1}{m}\right)^{(n-r)k}\right).$$

• As Z_i is an indicator, $E[Z_i] = Pr(Z_i = 1)$. Thus,

$$E[Z] = \sum_{i=0}^{m-1} E[Z_i] = m * E[Z_i]$$
$$= 2m \left(1 - \frac{1}{m}\right)^{nk} \left(1 - \left(1 - \frac{1}{m}\right)^{(n-r)k}\right).$$

For the leader election problem briefly introduced in Lecture Notes 15, we have n users, each with an identifier. Suppose that we have a good hash function (that looks uniform and independent), which outputs a b-bit hash value for each identifier. One way to solve the leader election problem is as follows: Each user obtains the hash value from its identifier, and the leader is the user with the smallest hash value. Give lower and upper bounds on the number of bits b necessary to ensure that a unique leader is successfully chosen with probability p. Make your bounds as tight as possible.

Assignment 3 - problem 4 (con't)

► Ans. (Lower Bound:) First, a unique leader is determined if and only if for some identifier *i*, it is mapped exactly by one user, while all other users are mapped to the identifiers larger than *i*. This implies that:

2^b-2 $p = \sum_{i=0}^{n} \Pr(\text{``identifier } i \text{ is mapped exactly by one user''})$ \cap "all other users are larger than i") $= \sum_{i=1}^{2^{\nu}-2} \frac{1}{2^{b}} \cdot \left(1 - \frac{i+1}{2^{b}}\right)^{n-1}$ (1) $\leq \sum_{i=0}^{2^{b}-2} rac{1}{2^{b}} \cdot \left(1 - rac{1}{2^{b}}
ight)^{n-1}$ $\leq \left(1-rac{1}{2^b}
ight)^{n-1}.$ Tutorial IV Slippers

▶ By re-arranging terms, we have:

$$b\geq \log_2\left(\frac{1}{1-p^{\frac{1}{n-1}}}\right).$$

Assignment 3 - problem 4 (con't)

- (Upper Bound:) Next, p is greater than the probability that user 1 is the unique leader. This happens when user 1 is mapped to uniquely to some number, and this number is smallest among all numbers mapped by other users. Thus,
 - $p \geq \Pr("user 1 is mapped uniquely" \cap "this number is smallest")$
 - = Pr("this number is smallest" | "user 1 is mapped uniquely") $\times Pr($ "user 1 is mapped uniquely")
 - = Pr("this number is smallest" | "user 1 is mapped uniquely")

$$\geq \frac{1}{n}\left(1-\frac{1}{2^{b}}\right)^{n-1}$$

By re-arranging terms, we have:

$$b \leq \log_2\left(rac{1}{1-(np)^{rac{1}{n-1}}}
ight)$$
Slippers Tutorial IV

► In conclusion,

$$\log_2\left(\frac{1}{1-p^{\frac{1}{n-1}}}\right) \leq b \leq \log_2\left(\frac{1}{1-(np)^{\frac{1}{n-1}}}\right).$$

- Consider an instance of SAT with *m* clauses, where every clause has exactly *k* literals.
 - 1. Give a Las Vegas algorithm that finds an assignment satisfying at least $m(1-2^{-k})$ clauses, and analyze its expected running time.
 - 2. Give a derandomization of the randomized algorithm using the method of conditional expectations.

- 1. Prove that, for every integer *n*, there exists a coloring of the edges of the complete graph K_n by two colors so that the total number of monochromatic copies of K_4 is at most $\binom{n}{4}2^{-5}$.
 - 2. Give a randomized algorithm for finding a coloring with at most $\binom{n}{4}2^{-5}$ monochromatic copies of K_4 that runs in expected time polynomial in n.
 - 3. Show how to construct such a coloring deterministically in polynomial time using the method of conditional expectations.

- hint:
 - 1. The expectation of the total number of monochromatic copies of K_4 .
 - 2. Las Vegas algorithm
 - 3. You can compare the expectation numbers of monochromatic copies of K_4 when a edge is assigned to different color.

Assignment 4 - problem 3

- Given an *n*-vertex undirected graph G = (V, E), consider the following method of generating an independent set. Given a permutation σ of the vertices, define a subset $S(\sigma)$ of the vertices as follows: for each vertex $i, i \in S(\sigma)$ if and only if no neighbor j of i precedes i in the permutation σ .
 - 1. Show that each $S(\sigma)$ is an independent set in G.
 - 2. Suggest a natural randomized algorithm to produce σ for which you can show that the expected cardinality of $S(\sigma)$ is

$$\sum_{i=1}^n \frac{1}{d_i+1}$$

where d_i denotes the degree of vertex *i*.

3. Prove that G has an independent set of size at least $\sum_{i=1}^{n} 1/(d_i + 1)$.

- Choose a random permutation σ from S_n, the set of all n! permutations of V.
- A vertex i is good if $\sigma(i) < \sigma(j)$ for every j adjacent to i.
- Let $S(\sigma)$ be the collection of good vertices and output $S(\sigma)$.

hint:

- 1. You can think about the definition of "good vertices", and proof it by contradiction.
- 2. The expected cardinality of $S(\sigma)$ is equal to the expected number of good vertices in G.

► We have shown using the probabilistic method that, if a graph G has n nodes and m edges, then there exists a partition of the n nodes into sets A and B such that at least m/2 edges cross the partition. Improve this result slightly: show that there exists a partition such that at least mn/(2n - 1) edges cross the partition.

hint:

- 1. You can fix a point and find the probability of the crossing edge.
- 2. You can discuss from two cases.
 - (a) n is odd.
 - (b) n is even.

▶ Use the general form of the Lovasz local lemma to prove that the symmetric version of Theorem 6.11 can be improved by replacing the condition $4dp \le 1$ by the weaker condition $ep(d+1) \le 1$.

hint:

- 1. Set $x_i = 1/(d+1)$ to the general case Lovasz local lemma.
- You have already had Pr(E_i) ≤ p from the symmetric version of Lovasz local lemma and you want to prove Pr(E_i) ≤ x_i ∏_{(i,j)∈E}(1 = x_j).